A Theoretical Study on the Orientation Problem in Linear Wireless Sensor Networks

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by

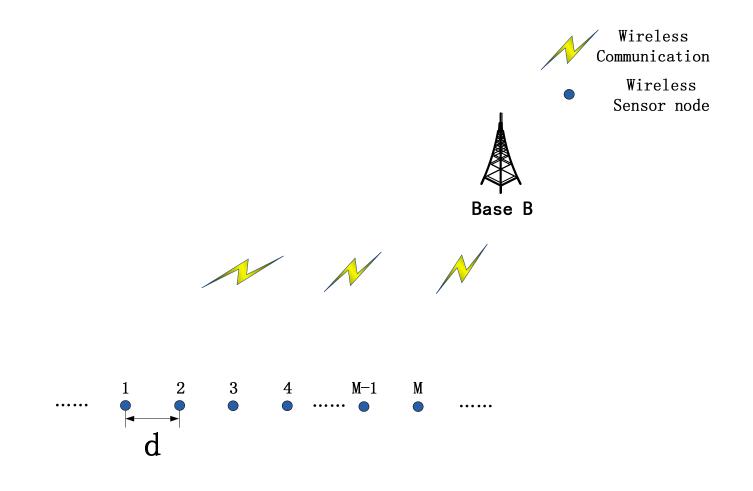
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What is Linear Sensor Networks

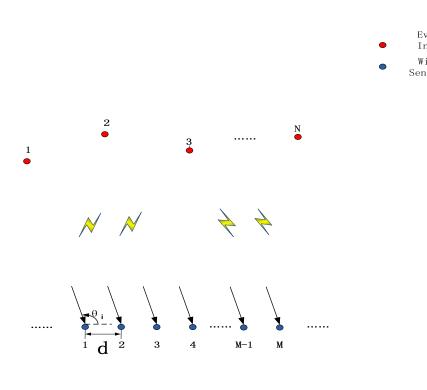
 Linear wireless sensor networks (LWSNs) are a special family of wireless sensor networks with regard to linear network topology. Compared with normal wireless sensor networks, LWSNs exhibit lower complexity. Even though simple in topology, LWSNs possess numerous practical applications, such as monitoring public transportations, oil pipes, factories and plants.

An Example of a LWSN



Model and Preconditions

The model of a LWSN



Suppose that at a time instance t,

Wireless ensor Node there are N events bursting out,
and they are sensed by M sensor nodes in the sensing area.

In the left figure, d denotes the distance between any two sensor nodes, and θ_i represents the angel of the i-th received signal with respect to the line of the deployment of M sensor nodes.

Model and Preconditions

Preconditions

- (1) Assume that all the sensor nodes are equally spaced, and the distance between any two sensor nodes is d;
- (2) Suppose that all the transmitted signals are all narrow-banded with some certain known frequency f;
- (3) All the signals propagate over the distance long enough such that the received signals are parallel to each other.

The Array Manifold Matrix

According to the previous network model and preconditions, the N received signals can be represented by an array manifold matrix

$$G = \begin{bmatrix} \frac{1}{2\pi d} \sin\theta_1 & \frac{1}{2\pi d} \sin\theta_2 & \cdots & \frac{1}{2\pi d} \sin\theta_1 \\ \frac{1}{\lambda} \sin\theta_1 & \exp\left(-\frac{j2\pi d}{\lambda} \sin\theta_2\right) & \cdots & \exp\left(-\frac{j2\pi d}{\lambda} \sin\theta_1\right) \\ \exp\left(-\frac{j(M-1)2\pi d}{\lambda} \sin\theta_1\right) & \exp\left(-\frac{j(M-1)2\pi d}{\lambda} \sin\theta_2\right) & \cdots & \exp\left(-\frac{j(M-1)2\pi d}{\lambda} \sin\theta_1\right) \end{bmatrix}_{\text{Maxim}}$$

In the above matrix,

(1)d represents the distance between sensor nodes;

- (2) $\lambda = 1/f$, where f is the frequency of the narrow-banded signals;
- (3) θ_i is the arriving angle of the i-th signal with respect to the line of the deployment of the M sensor nodes.

It is also assumed that these received N signals are interfered by some additive noise with zero mean and σ_n^2 Gaussian distribution which is uncorrelated with signals.

Therefore, the covariance matrix C of the received signal has the form of

$$C = C_{N} + GPG^{H}$$

where

- (1) C_N denotes the covariance matrix of M-dimension additive noise;
- (2) P denotes the correlation matrix between signals, and here, $P = diag(\sigma_1^2, \sigma_2^2, ..., \sigma_N^2) \cdot \sigma_i^2$ denotes the i-th signal power.

• Till now, it is clear that the covariance matrix can be parameterized by σ_n^2 , $\{\sigma_i^2\}$, and $\{\theta_i\}$, i.e.,

$$C = C(\sigma_n^2, {\sigma_i^2}, {\theta_i})$$

it is obvious that the covariance matrix conveys knowledge about the arrival angle of the i-th signal in that C is a function of which will help us determine the direction of this signal.

 The structure of the covariance matrix The covariance matrix C is a complex valued toeplitz matrix, and also a Hermitian matrix. However, it is still not easy to estimate this matrix even though it is quite structured. Therefore, we firstly introduce the method which can be used to estimate some structured real valued matrices, and then the result will be extended to the realm of complex valued matrices.

The Estimation of real valued covariance matrix

If all the signals are real valued, we want to estimate the covariance matrix of signals from N samples $x_1, x_2, ..., x_N$. Here, we suppose that the N samples are drawn from an M-dimension Gaussian distribution with zero mean and covariance matrix C. Therefore, he probability density function for each x_i is

$$p(x_i) = (2\pi)^{-\frac{M}{2}} \det(C)^{-\frac{1}{2}} \exp(-\frac{x_i^T C^{-1} x_i}{2})$$

What is more, If all these samples are independent to each other, the joint probability density function of these N samples is

$$p(x_1, x_2, \dots, x_N) = (2\pi)^{-\frac{MN}{2}} \det(C)^{-\frac{N}{2}} \exp(-\sum_{i=1}^{N} \frac{x_i^T C^{-1} x_i}{2})$$

The log-likelihood function of C is

$$\ln\left(L\left(C;\left\{x_{i}\right\}\right)\right) = \frac{MN}{2}\ln 2\pi - \frac{N}{2}\ln \det\left(C\right) - \frac{1}{2}tr(C^{-1}\sum_{i=1}^{N}x_{i}x_{i}^{T})$$

Maximum-likelihood estimation is used here, hence,

$$\begin{split} \hat{C} &= \arg\max(L\,{}^{\shortmid}\!(C; \{x_i\})) \\ &= \arg\max(-\ln\det\left(C\right) - tr(C^{-1}\,\frac{1}{N}\sum_{i=1}^N x_i x_i^T)) \end{split}$$

The matrix $S = \frac{1}{N} \sum_{i=1}^{N} x_i x_i^T$ is of course a semipositive definite matrix, and is also symmetric.

If we want to acquire the estimation of \mathcal{C} , the following optimization problem should be solved:

- maximize: $-\ln \det \left(C\right) tr(C^{-1}\frac{1}{N}\sum_{i=1}^N x_i x_i^T)$ • subject to: C belongs to the set of semi-positive
- subject to: C belongs to the set of semi-positive definite matrices

It is proved that the above optimization problem will have solutions.

How to solve the estimation problem
 If the objective function

-
$$\ln \det (C)$$
 - $tr(C^{-1}S)$

achieves its maximum, he total differential of this function will be zero, i.e.,

$$d\left(-\ln\det\left(C\right) - tr(C^{-1}S)\right) = 0$$

Here, we use two facts:

$$\textbf{(1)} \ d\Big(\det \Big(C\Big)\Big) = \sum_{i=1}^{M} \frac{\partial \sum_{j=1}^{M} c_{ij} C_{ij}^{*}}{\partial c_{ij}} dc_{ij} = tr(C^{-1}dC)$$

(2)
$$d(tr(C^{(-1)}S)) = -tr(C^{(-1)}dCC^{(-1)}S)$$

where C^* is the adjoint matrix of C.

Hence,

$$d\left(-\ln \det\left(C\right) - tr(C^{-1}S)\right) = tr(\left(C^{-1}SC^{-1} - C^{-1}\right)dC)$$

If C is symmetric, we can substitute dC by C, therefore,

$$d\left(-\ln\det\left(C\right) - tr(C^{-1}S)\right) = 0$$



$$tr(C^{-1}S) = M$$

where M is the dimension of x_i . Here, the estimator of C is $S = \frac{1}{N} \sum_{i=1}^{N} x_i x_i^T$.

Estimation of a complex valued covariance matrix

In previous sections, when the covariance matrix is real valued, the best estimator of it is the average of the sample matrix. However, the covariance of the received signals has the form of

$$C = C_{N} + GPG^{H}$$

where G is the array manifold matrix which is definitely a complex valued matrix.

The received signal x_i is complex valued, so we define

$$S = \frac{1}{N} \sum_{i=1}^{N} x_i x_i^H$$

where H denotes the conjugate transpose.

Solving estimation problem of complex valued covariance matrix

We can use the result of previous sections to solve the estimation problem of complex valued scenario.

The key point to this problem is to construct a real valued, positive definite, and symmetric matrix using S.

Let
$$S^R = \frac{1}{N} \sum_{i=1}^{N} \begin{bmatrix} \operatorname{Re}(x_i) \\ \operatorname{Im}(x_i) \end{bmatrix} [\operatorname{Re}(x_i)^T \quad \operatorname{Im}(x_i)^T]$$

Obviously, S^R is a real valued, positive definite, and symmetric matrix, and it conveys all the information about S.

Furthermore, let

$$O = \frac{1}{N} \sum_{i=1}^{N} \operatorname{Re}(x_{i}) \operatorname{Re}(x_{i})^{T}$$

$$P = \frac{1}{N} \sum_{i=1}^{N} \operatorname{Im}(x_{i}) \operatorname{Im}(x_{i})^{T}$$

$$Q = \frac{1}{N} \sum_{i=1}^{N} \operatorname{Re}(x_{i}) \operatorname{Im}(x_{i})^{T}$$

then

$$S^R = \begin{bmatrix} O & Q \\ Q^T & P \end{bmatrix}$$

Denote the estimator acquired from S^R is \hat{C}^R .

The relationship between \hat{C}^R and \hat{C} is one-to-one mapping, i.e., .

$$\hat{c}_{ij} = o_{ij} + p_{ij} - jq_{ij} + jq_{ji}$$

Therefore, the estimation problem is solved through an estimation of real valued covariance matrix.

Future Work

This paper discuss on the theoretical parts of determining the arrival angles of sensed events using the covariance matrix estimation. In future, intensive field experiment should be conducted to verify the result of this paper.

Thank you!